# Original research

# SARS-CoV-2 infection and adverse outcomes in users of ACE inhibitors and angiotensin-receptor blockers: a nationwide case-control and cohort analysis

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# ABSTRACT

**Objective** To examine the impact of ACE inhibitor (ACE-I)/angiotensin receptor blocker (ARB) use on rate of SARS-CoV-2 infection and adverse outcomes.

**Methods** This nationwide case-control and cohort study included all individuals in Denmark tested for SARS-CoV-2 RNA with PCR from 27 February 2020 to 26 July 2020. We estimated confounder-adjusted ORs for a positive test among all SARS-CoV-2 tested, and inverse probability of treatment weighted 30-day risk and risk ratios (RRs) of hospitalisation, intensive care unit (ICU) admission and mortality comparing current ACE-I/ARB use with calcium channel blocker (CCB) use and with non-use.

**Results** The study included 13 501 SARS-CoV-2 PCR-positive and 1 088 695 PCR-negative individuals. Users of ACE-I/ARB had a marginally increased rate of a positive PCR when compared with CCB users (aOR 1.17, 95% CI 1.00 to 1.37), but not when compared with nonusers (aOR 1.00 95% CI 0.92 to 1.09).

Among PCR-positive individuals, 1466 (11%) were ACE-I/ARB users. The weighted risk of hospitalisation was 36.5% in ACE-I/ARB users and 43.3% in CCB users (RR 0.84, 95% CI 0.70 to 1.02). The risk of ICU admission was 6.3% in ACE-I/ARB users and 5.4% in CCB users (RR 1.17, 95% CI 0.64 to 2.16), while the 30-day mortality was 12.3% in ACE-I/ARB users and 13.9% in CCB users (RR 0.89, 95% CI 0.61 to 1.30). The associations were similar when ACE-I/ARB users were compared with non-users.

**Conclusions** ACE-I/ARB use was associated neither with a consistently increased rate nor with adverse outcomes of SARS-CoV-2 infection. Our findings support the current recommendation of continuing use of ACE-Is/ ARBs during the SARS-CoV-2 pandemic. **Trial registration number** EUPAS34887

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## INTRODUCTION

The use of renin-angiotensin system inhibitors, including ACE inhibitors (ACE-Is) and angiotensin receptor blockers (ARBs), has been suggested to increase the risk of being infected by SARS-CoV-2 and of adverse outcomes of coronavirus disease

## Key messages

## What is the key question?

Does use of ACE inhibitors (ACE-I) and angiotensin receptor blockers (ARBs) increase the rate or worsen the outcome of SARS-CoV-2 infection?

### What is the bottom line?

 Use of ACE-I/ARB was associated neither with a consistently increased rate nor with adverse outcomes of SARS-CoV-2 infection, compared with ACE-I/ARB non-use or calcium channel blocker use.

## Why read on?

This population-based study, including all individuals PCR tested for SARS-CoV-2 in Denmark, used extensive confounder adjustment and an active comparator design to examine the association between ACE-I/ARB use and the rate of microbiologically verified SARS-CoV-2 infection as well as the associated outcomes including hospitalisation, intensive care unit admission, mechanical ventilation and death.

2019 (COVID-19) caused by SARS-CoV-2. ACE-Is/ ARBs upregulates the human ACE2 receptor, which facilitates entry of SARS-CoV-2 into cells.<sup>1-5</sup> The initially published studies of ACE-I/ARB users and SARS-CoV-2 infection reported no increased risk or worsened outcome after a positive SARS-CoV-2 test or a diagnosis of COVID-19.6-16 However, available studies are limited by incomplete data on recent ACE-I/ARB use and preexisting comorbidities,<sup>679</sup> by restriction to hospitalised or hospitaldiagnosed COVID-19 patients,<sup>9 10 12-15</sup> or by incomplete follow-up.<sup>6</sup> Other limitations have included immortal time bias from inclusion of in-hospital ACE-I/ARB use after COVID-19 diagnosis,<sup>10 11 15 16</sup> which may lead to apparently beneficial effects, because ACE-I/ARBs would only be prescribed if patients are haemodynamically stable

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and have survived until drug initiation. Only few studies on risk of SARS-CoV-2 infection included an active comparator.<sup>12 13</sup>

Major professional societies have called for further highquality research and issued warnings against ACE-I/ARB discontinuation in patients with SARS-CoV-2 infection,<sup>1 17</sup> to avoid worsening of underlying cardiometabolic conditions.<sup>18</sup>

As ACE-Is/ARBs are widely used drugs, any association with rate of infection or adverse outcomes of SARS-CoV-2 infection may have an important public health impact. Thus, there is an urgent need to examine the hypotheses of increased rate of SARS-CoV-2 infection and worsened outcome, among users of ACE-Is/ARBs.

## METHODS

## Design and setting

We conducted this nationwide combined case-control and cohort study in Denmark (population  $\sim$ 5.8 million persons) to study the rate (case-control design) and the prognosis (cohort design) of SARS-CoV-2. Denmark has a uniform tax-supported healthcare system responsible for all acute care in Denmark. The unique personal identification number assigned at birth or on immigration allows individual-level linkage of data and complete follow-up.<sup>19</sup>

As previously described,<sup>20</sup> data on all individuals tested for SARS-CoV-2 RNA by reverse transcriptase PCR (RT-PCR) were obtained from the Danish Microbiology Database<sup>21</sup> and linked to regularly updated data from the Danish National Patient Registry,<sup>22</sup> the Civil Registration System,<sup>23</sup> the National Prescription Database<sup>24</sup> and the Danish Register of Causes of Deaths.<sup>25</sup>

#### Participants

The study included prospectively collected data on all individuals tested for SARS-CoV-2 RNA from 27 February 2020 (the date when the first patient tested positive for SARS-CoV-2 in Denmark) to 26 July 2020, allowing complete 30-day follow-up through 26 August 2020.

The study population for the analysis of the rate of acquiring a positive SARS-CoV-2 test (the case-control analysis) included all patients tested for SARS-CoV-2, while the analysis of the outcome (the cohort analysis) only included patients who tested positive for SARS-CoV-2. The first date of a positive test was defined as the index date.

### **Medication exposures**

Current use of ACE-Is and ARBs was ascertained from prescriptions filled within 90 days before testing (online supplemental eFigure 1). Former use was defined as a prescription filled within 91–365 days before testing, while non-use was defined as no prescriptions filled during 1 year before testing. Prescription data included complete and valid information on all prescriptions filled at community pharmacies in Denmark since 1995.<sup>24</sup> Registry data include date of dispensing, Anatomical Therapeutic Chemical code and drug quantity. In addition to prescriptions for ACE-I/ARBs, we obtained data on prescriptions for other antihypertensive medications, including calcium channel blockers (CCBs), beta blockers and thiazides.

The main exposure comparisons were current use of ACE-Is/ ARBs versus current use of CCBs, and current use of ACE-Is/ ARBs versus no ACE-Is/ARB use. The first comparison allowed us to reduce confounding by indication, as CCB is an active drug with medical indications similar to those for ACE-Is/ARBs and with no known effects on the renin–angiotensin system.

## SARS-CoV-2 infection and adverse outcomes

Cases in the case-control analysis were individuals acquiring a positive SARS-CoV-2 PCR test, and controls were individuals with a negative PCR test. The primary outcome in the cohort analysis of outcome was death within 30 days following a positive SARS-CoV-2 test. Secondary outcomes included hospital admission (hospital stay lasting >12 hours) within 30 days after index date, among patients who were not already hospitalised. Additional secondary outcomes included intensive care unit (ICU) admission, ICU admission with mechanical ventilation (MV) and dialysis within 30 days after index date. Secondary outcomes were obtained from the Danish National Patient Registry, which is also the platform for mandatory reporting to the national database for quality of ICU, and earlier assessments of the accuracy of its data on ICU admissions, ICU admission with MV and dialysis yielded positive predictive values of 96%-100%.<sup>26 27</sup>

### **Potential confounders**

We included a wide range of potential confounders that may be associated with both ACE-I/ARB use and the risk and outcome of SARS-CoV-2 infection (table 1). Data on age, sex, marital status, ethnicity and urban residence were obtained from the Civil Registration System.<sup>23</sup> We also obtained information on comorbidities associated with an inpatient stay or an outpatient hospital clinic diagnosis or with treatment by prescribed medications within 10 years before the index date (codes provided in online supplemental eTable 1).<sup>22 24</sup> Prescriptions for concurrent medications filled within 90 days before the index date were also included.<sup>24</sup>

### Statistical methods

In the test-negative case-control analysis, the association between ACE-I/ARB use and a positive SARS-CoV-2 test was examined using a logistic regression model to compute ORs with 95% CIs adjusted for all covariates in table 1 and test date in 3-day intervals. In post hoc analyses, we tabulated patient characteristics in all tested patients by exposure group (current ACE-I/ARB use, no ACE-I/ARB use and current CCB use) and described positive:negative ratio, both overall and stratified by age group.

In the cohort analyses of outcome, we computed a propensity score (PS) for each individual, that is, the probability of being exposed in each comparison (eg, current ACE-I/ARB use vs current CCB use), using a logistic regression model including all covariates in table 1. We used the PS for inverse probability of treatment weighting (IPTW) with stabilised weights to estimate the average treatment effect in the population.<sup>28</sup> Covariate balance before and after IPTW was described using standardised mean differences.

We restricted the comparison of current vs no ACE-I/ARB use to patients aged 50 years or older as there were only few ACE-I/ ARB users, but many non-users, under age 50 leading to imbalance before restriction.

The cohort analysis of adverse outcomes followed patients from the day of their first positive test result to the date of the outcome of interest, that is, hospital admission, ICU admission (with or without need for MV), dialysis, date of death, emigration, or for up to 30 days. Patients were excluded in the analysis of each non-fatal outcome if they experienced that outcome from seven to 1 day before the test date.

Patient characteristics were tabulated according to each exposure group before and after weighting. For each exposure group, we estimated the 30-day weighted risks and risk differences with

Table 1	Characteristics of patients with positive and negative tests
for SARS-	CoV-2

for SARS-Cov-2	Test-positive	Test-negative
	cases	controls
	(n=13501)	(n=1 088 695)
Age, median (IQR)	47 (31–61)	43 (26–59)
Sex (male)	5763 (42.7)	474335 (43.6)
Status as healthcare professional		
Physician	423 (3.1)	10247 (0.9)
Nurse	1351 (10.0)	31 520 (2.9)
Care assistant	607 (4.5)	24211 (2.2)
Use of antihypertensives (within prior 90 days)		
ACE inhibitors	587 (4.3)	46 195 (4.2)
Angiotensin receptor blockers	887 (6.6)	63 467 (5.8)
Calcium channel blockers	690 (5.1)	55 901 (5.1)
Thiazides	646 (4.8)	47 627 (4.4)
Beta blockers	833 (6.2)	60 685 (5.6)
Other antihypertensives	47 (0.3)	3336 (0.3)
Diagnoses (within prior 10 years)		
Hypertension	1416 (10.5)	99 572 (9.1)
Atrial fibrillation	652 (4.8)	40 165 (3.7)
Hospital-diagnosed obesity	931 (6.9)	71 682 (6.6)
Angina	610 (4.5)	42 965 (3.9)
Heart valve disease	272 (2.0)	18917 (1.7)
Alcohol abuse	224 (1.7)	27 855 (2.6)
Diabetes	1162 (8.6)	69 796 (6.4)
Venous thromboembolism	296 (2.2)	16 947 (1.6)
Dementia	384 (2.8)	9980 (0.9)
Myocardial infarction	211 (1.6)	14487 (1.3)
Liver disease	124 (0.9)	10 449 (1.0)
Kidney disease	291 (2.2)	15 955 (1.5)
Chronic pulmonary disease	908 (6.7)	73 435 (6.7)
End-stage renal disease	21 (0.2)	1215 (0.1)
Heart failure	292 (2.2)	17515 (1.6)
Cancer	737 (5.5)	61 394 (5.6)
Stroke	413 (3.1)	23500 (2.2)
Markers of smoking	3293 (24.4)	296 588 (27.2)
Medications (within prior 90 days)		
Statins	1244 (9.2)	94770 (8.7)
Antivirals	193 (1.4)	12171 (1.1)
Low-dose aspirin	476 (3.5)	36812 (3.4)
Vitamin K-antagonists	106 (0.8)	8138 (0.7)
Opioids	778 (5.8)	51 647 (4.7)
Antidepressants	934 (6.9)	74239 (6.8)
Immunosuppressants	47 (0.3)	4207 (0.4)
Antipsychotics	262 (1.9)	22 504 (2.1)
Glucocorticoids	278 (2.1)	19675 (1.8)
Hypnotics	430 (3.2)	29887 (2.7)
Loop diuretics	524 (3.9)	29630 (2.7)
Antibiotics	2042 (15.1)	111 921 (10.3)
		Continued

Table 1   Continued		
	Test-positive cases (n=13 501)	Test-negative controls (n=1 088 695)
Socioeconomic factors		
Marital status		
Widowed	1005 (7.4)	53203 (4.9)
Divorced	1568 (11.6)	121 025 (11.1)
Married	6205 (46.0)	429375 (39.4)
Unmarried	4653 (34.5)	483 512 (44.4)
Unknown	70 (0.5)	1580 (0.1)
Ethnicity		
Non-immigrant	10329 (76.5)	935 629 (85.9)
First-generation or second- generation immigrant	3172 (23.5)	153066 (14.1)
Residence		
Non-urban	6539 (48.4)	648923 (59.6)
Urban	6962 (51.6)	439772 (40.4)

robust 95% CIs using generalised linear models with a binomial distribution and an identity link. Risk ratios (RRs) were estimated similarly but using a log link.

A subgroup analysis was conducted to address potential effect modification by age group ( $\leq 65$  years, > 65 years). A second subgroup analysis was conducted in patients with assumed uncomplicated hypertension as the primary indication for treatment (defined as patients without a history of diabetes, renal disease, angina pectoris, myocardial infarction or heart failure). A third subgroup analysis was restricted to patients tested after the test strategy changed in Denmark and the country was locked down on 13 March 2020.

In a sensitivity analysis, we explored the robustness of our findings by repeating the analyses for patients who filled a prescription within 120 days before a positive test. We also conducted a sensitivity analysis stratified by calendar time to address changes in testing and hospitalisation strategy. We followed the protocol registered in the European Union electronic Register of Post-Authorisation Studies (EU PAS Register) (record number: EUPAS34887), with two exemptions. First, standardised mortality rate weighting was replaced with IPTW with stabilised weights to better handle the low number of current CCB users in the comparison group. Second, the comparison with non-users of ace-i/ARB was restricted to patients aged 50 years and older due to the imbalance mentioned above.

All data management and statistical analyses were performed using STATA V.16 MP.

## RESULTS

The study included 13501 test-positive and 1088695 testnegative individuals (table 1). Among test-positive cases, there were 1466 current ACE-I/ARB users (11%) (587 ACE-I users, 887 ARB users and 8 using both) of whom 1336 were 50 years or above and included in the comparison between current ACE-I/ ARB use with no ACE-I/ARB use. Among the 1466 current ACE-I/ARB users, 1065 did not receive concurrent CCBs and could therefore be included in the comparison between current ACE-I/ARB use and current CCB use.

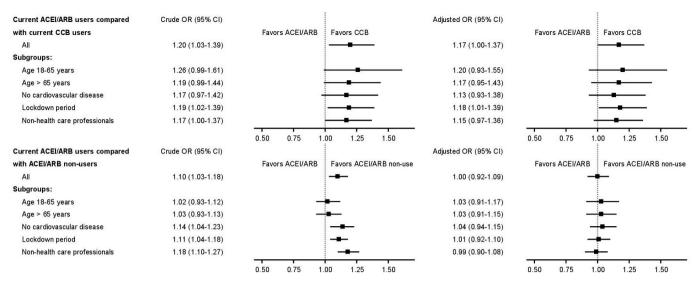


Figure 1 OR for a positive test among ACE-I/ARB users compared with CCB users and non-users tested for SARS-CoV2. ACE-I, ACE inhibitors; ARB, angiotensin receptor blocker; CCB, calcium channel blocker.

### Rate of acquiring a positive SARS-CoV-2 test

There were no major differences in characteristics of test-positive and test-negative individuals with regard to age, sex, comorbidity and medication use (table 1). There were, however, more healthcare professionals, more living in urban areas, but fewer non-immigrants among the test-positive cases. After adjusting, ACE-I/ARB users had a slightly increased rate of a positive test compared with current CCB users (adjusted OR 1.17, 95% CI 1.00 to 1.37) (figure 1). ACE-I/ARB users had a similar rate of a positive test compared with ACE-I/ARB non-users (adjusted OR 1.00, 95% CI 0.92 to 1.09) (figure 1).

#### **Outcomes of SARS-CoV-2 infection**

Patient characteristics of current ACE-I/ARB users are compared with CCB users in table 2 and with non-users in table 3.

After excluding patients both on ACE-I/ARB and CCB from the SARS-CoV-2 positive cohort, 1065 ACE-I/ARB users and 289 current CCB users remained. The characteristics of patients in the two groups were well balanced after weighting (table 2).

The risk of hospital admission was 36.5% in ACE-I/ARB users and 43.3% in CCB users (RR 0.84, 95% CI 0.70 to 1.02) after weighting (table 4). The risk of ICU admission was 6.3% in ACE-I/ARB users and 5.4% in CCB users (RR 1.17, 95% CI 0.64 to 2.16), while the risk of death within 30 days was 12.3% in ACE-I/ARB users and 13.9% in CCB users (RR 0.89, 95% CI 0.61 to 1.30).

Among SARS-CoV-2 positive patients aged 50 years or above, 1336 were current ace-i/ARB users and 4326 were non-users. Characteristics of patients in the two groups were well balanced after weighting (table 3). The weighted risk of hospital admission was 32.3% in ACE-I/ARB users and 33.1% in non-users (RR 0.98, 95% CI 0.84 to 1.14) (table 4). The risk of ICU admission was 6.4% in ACE-I/ARB users and 5.1% in non-users (RR 1.25, 95% CI 0.66 to 2.36), while the risk of death within 30 days was 9.5% in ACE-I/ARB users and 10.8% in non-users (RR 0.88, 95% CI 0.68 to 1.14).

### Subgroup analyses

Results were confirmed in subgroup analyses stratified by age, no cardiovascular disease other than hypertension, and period of testing (online supplemental eTable 2). Results for various exposure definitions are shown in online supplemental eTable 3. Results remained robust when the exposure window was changed to 120 days and when current ACE-I/ ARB use was compared with current thiazide and beta-blocker use. There was no conclusive difference in the associations when analysing current ACE-I and ARB use separately, although point estimates differed.

#### Post hoc analyses

Among all tested individuals, patient characteristics were similar for current ACE-I/ARB users compared with current CCB users, while ACE-I/ARB non-users were younger and had less comorbidity (online supplemental eTable 4). The overall positive:negative ratios were similar for current ACE-I/ARB users, ACE-I/ARB non-users and current CCB users (online supplemental eTable 5). When stratified by age, the positive:negative ratio differed slightly between the exposure groups among young adults only.

### DISCUSSION

In this large nationwide study of prospectively collected data, current ACE-I/ARB use was associated with a marginally increased rate of testing positive compared with CCB use but not when compared with ACE-I/ARB non-use. Moreover, we showed that among individuals who tested positive for SARS-CoV-2, current use of ACE-Is/ARBs was not associated with increased risk of hospitalisation, ICU admission or 30-day mortality.

### Strengths and limitations of the study

Strengths of the current study include its nationwide populationbased design, including all individuals tested for SARS-CoV-2 in Denmark, with linkage to validated medical registries and complete follow-up for censoring of all Danish residents.<sup>19 20 23-25</sup> ICU admissions and treatments also are recorded accurately, as the Danish National Patient Registry is used for financial reimbursements to hospitals and for mandatory reporting to national quality of care databases.<sup>26 27</sup> The RT-PCR test for presence of SARS-CoV-2 has high sensitivity and predictive value,<sup>29</sup> and registration of the tests are complete and accurately recorded in the Danish Microbiology Database.<sup>21</sup> Despite these strengths, the study had several limitations. The threshold for SARS-CoV-2

	Unweighted cohort	t		IPTW cohort		
	Current ACE-I/ARB users Current CCB users			Current ACE-I/ARB users	E-I/ARB Current CCB users	
	(n=1065)	(n=289)	SMD	(n=1065)	(n=285)	SMD
Age, median (IQR)	68 (58–79)	72 (60–82)	0.20	69 (58–80)	68 (57–80)	0.02
Sex (male)	519 (48.7)	133 (46.0)	0.05	511 (48.0)	139 (48.9)	0.02
Status as healthcare professional						
Physician	20 (1.9)	(n<5)	0.07	18 (1.7)	(n<5)	0.00
Nurse	46 (4.3)	5 (1.7)	0.15	40 (3.8)	10 (3.3)	0.02
Care assistant	43 (4.0)	9 (3.1)	0.05	41 (3.9)	11 (4.0)	0.01
Use of antihypertensives (within prior 90 days)						
ACE inhibitors	426 (40.0)	0	1.15	426 (40.1)	0	1.16
Angiotensin receptor blockers	645 (60.6)	0	1.75	644 (60.5)	0	1.75
Calcium channel blockers	0	289 (100.0)		0	285 (100.0)	
Thiazides	316 (29.7)	34 (11.8)	0.45	275 (25.9)	72 (25.3)	0.01
Beta blockers	234 (22.0)	70 (24.2)	0.05	239 (22.4)	65 (22.6)	0.01
Other antihypertensives	16 (1.5)	11 (3.8)	0.14	22 (2.1)	6 (2.2)	0.01
Diagnoses (within prior 10 years)						
Hypertension	425 (39.9)	137 (47.4)	0.15	440 (41.3)	122 (42.6)	0.03
Atrial fibrillation	138 (13.0)	37 (12.8)	0.00	138 (13.0)	38 (13.4)	0.01
Hospital-diagnosed obesity	112 (10.5)	23 (8.0)	0.09	106 (10.0)	26 (9.1)	0.03
Angina	146 (13.7)	33 (11.4)	0.07	142 (13.3)	41 (14.2)	0.03
Heart valve disease	69 (6.5)	13 (4.5)	0.09	64 (6.0)	13 (4.5)	0.07
Alcohol abuse	35 (3.3)	12 (4.2)	0.05	37 (3.5)	11 (3.8)	0.02
Diabetes	291 (27.3)	59 (20.4)	0.16	276 (26.0)	78 (27.3)	0.03
Venous thromboembolism	45 (4.2)	22 (7.6)	0.14	52 (4.8)	13 (4.4)	0.02
Dementia	66 (6.2)	20 (6.9)	0.03	67 (6.3)	16 (5.7)	0.03
Myocardial infarction	60 (5.6)	8 (2.8)	0.14	53 (5.0)	11 (4.0)	0.05
Liver disease	10 (0.9)	5 (1.7)	0.07	13 (1.2)	(n<5)	0.00
Kidney disease	66 (6.2)	37 (12.8)	0.23	81 (7.6)	22 (7.6)	0.00
Chronic pulmonary disease	128 (12.0)	34 (11.8)	0.01	127 (11.9)	34 (11.9)	0.00
End-stage renal disease	(n<5)	6 (2.1)	0.18	6 (0.5)	(n<5)	0.01
Heart failure	111 (10.4)	15 (5.2)	0.20	99 (9.3)	24 (8.6)	0.03
Cancer	136 (12.8)	32 (11.1)	0.05	132 (12.4)	33 (11.6)	0.02
Stroke	98 (9.2)	35 (12.1)	0.09	104 (9.8)	26 (9.1)	0.02
Markers of smoking	354 (33.2)	93 (32.2)	0.09	351 (33.0)	88 (30.7)	0.02
Markers of smoking Medications (within prior 90 days)	55- (55.2)	JJ (JL.L)	0.02	55: (55.0)	00 (00.7)	0.00
Statins	381 (35.8)	95 (32.9)	0.06	373 (35.1)	108 (37.7)	0.06
Antiviral	14 (1.3)	(n<5)	0.11	12 (1.1)	(n<5)	0.05
Low-dose aspirin	14 (1.3)	(1<5) 37 (12.8)	0.01	140 (13.1)	(1<5) 37 (12.8)	0.05
Vitamin K-antagonists	23 (2.2)	9 (3.1)	0.01	25 (2.4)	6 (2.0)	0.01
Opioids	131 (12.3)	9 (3.1) 46 (15.9)	0.06	141 (13.3)	6 (2.0) 45 (15.7)	0.03
Antidepressants			0.06			0.07
•	152 (14.3)	47 (16.3)	0.06	156 (14.6)	42 (14.6)	0.00
Immunosuppressants	8 (0.8)	(n<5)		9 (0.9)	(n<5)	
Antipsychotics	42 (3.9)	10 (3.5)	0.03	41 (3.8)	12 (4.1)	0.01
Glucocorticoids	45 (4.2)	11 (3.8)	0.02	43 (4.1)	9 (3.0)	0.06
Hypnotics	71 (6.7)	20 (6.9)	0.01	75 (7.1)	26 (9.1)	0.08
Loop diuretics	157 (14.7)	42 (14.5)	0.01	158 (14.9)	45 (15.7)	0.02
Antibiotics Socioeconomic factors	282 (26.5)	80 (27.7)	0.03	283 (26.6)	77 (27.1)	0.01

Continued

# Table 2 Continued

	Unweighted cohort			IPTW cohort			
	Current ACE-I/A users	ARB Current CCB users		Current ACE-I/ARB users	Current CCB users	5	
	(n=1065)	(n=289)	SMD	(n=1065)	(n=285)	SMD	
Marital status							
Widowed	207 (19.4)	72 (24.9)	0.13	221 (20.8)	63 (22.1)	0.03	
Divorced	156 (14.6)	34 (11.8)	0.09	149 (14.0)	35 (12.2)	0.05	
Married	589 (55.3)	156 (54.0)	0.03	584 (54.9)	158 (55.2)	0.01	
Unmarried	102 (9.6)	24 (8.3)	0.04	100 (9.4)	28 (9.8)	0.01	
Unknown	11 (1.0)	(n<5)	0.00	11 (1.0)	(n<5)	0.02	
Ethnicity							
Non-immigrant	902 (84.7)	251 (86.9)	0.06	908 (85.3)	240 (84.1)	0.03	
First- or second-generation immigrant	163 (15.3)	38 (13.1)	0.06	156 (14.7)	45 (15.9)	0.03	
Residence							
Non-urban	581 (54.6)	160 (55.4)	0.02	584 (54.9)	155 (54.4)	0.01	
Urban	484 (45.4)	129 (44.6)	0.02	480 (45.1)	130 (45.6)	0.01	

ACE-I, ACE inhibitors; ARB, angiotensin receptor blocker; CCB, calcium channel blocker; IPTW, inverse probability of treatment weighting; SMD, standardised mean difference.

testing was lowered as the pandemic progressed and testing capacity has increased steadily since March 2020.<sup>20</sup> Importantly, our results did not change after restriction to patients tested after the policy guidelines changed and the country was locked down on 13 March 2020. Any bias should be minimal as use of ACE-I/ARB was not expected to influence the threshold for testing, which followed national guidelines. This was confirmed by the lack of major difference in positive:negative ratios between exposure groups.

Medication use was assessed using complete and valid data on prescriptions prior to testing in a time window corresponding to the typical interval between medication refills. Still, misclassification could have occurred if some patients had sporadic use of drugs filled more than 90 days before testing. Given the chronic use of the drugs included in the study, such misclassification should be minor and not associated with the outcome of interest. The direction of any information bias is therefore expected to be towards the null. It is likely to be minimal as no change in estimates was observed after the time window was extended from 90-120 days before testing in a sensitivity analysis. A related concern is that we lacked in-hospital medication data to examine the impact of continued and discontinued use of prescribed drugs during follow-up, which may be relevant, since 50% of hospitalised COVID-19 patients may discontinue ACE-I/ARB.30

The use of ICU admission as an outcome in observational prognostic studies is challenging.<sup>31</sup> In clinical practice, ICU admission is offered to patients who are expected to have a clear prognostic benefit from invasive monitoring and treatment.<sup>32</sup> This is the case particularly in countries with limited ICU capacity, which are not able to accommodate the level of need during the COVID-19 pandemic. In Denmark, ICU capacity was more than sufficient during the study period. In addition, patients' quality of life, functional level at home and hospital capacity may influence the decision to admit a patient to the ICU. This may explain why ICU and mortality outcomes tended to go in opposite directions for some associations examined in our study.

Potential confounding by medical indication for drug treatment was handled by use of an active comparator in the main analysis, by implementing IPTW that included a large number of potential confounders, and by restricting analyses to subgroups according to indication for treatment. Although cardiovascular and other diagnoses used in the study have documented high positive predictive values,<sup>20 33</sup> we cannot entirely rule out that our findings were influenced by unmeasured confounding by indication and contraindication for treatments and by severity of underlying comorbidity. A final concern is that precision of risk estimates was limited in some subgroups.

# Comparison with other studies

Our results on risk of acquiring a positive SARS-CoV-2 test are consistent with the few published studies on risk of ACE-I/ARB users compared with non-users, <sup>6–8</sup> <sup>12</sup> <sup>13</sup> <sup>16</sup> and to other antihypertensives. <sup>12</sup> <sup>13</sup> The earlier studies, which used hospital based rather than population based data or were limited by incomplete data on recent ACE-I/ARB exposure, reported ORs ranging from 0.94 to 1.23 associated with ACE-I/ARB use. <sup>6–8</sup> <sup>12</sup> <sup>13</sup> <sup>16</sup> A Danish study limited to hospital-diagnosed COVID-19 cases and population controls found similar association between ACE-I/ARB use and non-use (HR 1.05, 95% CI 0.80 to 1.36) and CCB use (HR 1.23, 95% CI 0.89 to 1.70), respectively. <sup>13</sup>

Thus, our study has confirmed and extended the previous findings in a large nationwide cohort of all individuals with microbiologically confirmed SARS-CoV-2 infection comparing ACE-I/ARB users to a specific active comparator group of users of another antihypertensive, CCB, corroborating no major association.

In our cohort analysis of patients with positive SARS-CoV-2 test, we found that ACE-I/ARB users had no increased risk of an ICU admission, confirming previous hospital-based studies.<sup>9-11</sup> <sup>13-15</sup> More importantly, our findings qualified that ACE-I/ARB use was not associated with increased risk of hospitalisation when accounting for confounding factors. Our adverse outcome findings remained robust across varying exposure definitions, subgroups and also in comparison with CCB use.

Our observed associations were more precise and closer to the null than reported in a recent systematic review.<sup>34</sup> The few larger studies had some limitations. Reynolds *et al* conducted a study of 12594 patients tested for SARS-CoV-2 in emergency

 Table 3
 Characteristics of current ACE-I/ARB users and non-users aged 50 years or above, before and after inverse probability of treatment weighting

	Unweighted coh	Unweighted cohort			IPTW cohort		
	Current ACE-I/ARB	Current ACE-I/ARB use No ACE-I/ARB use		Current ACE-I/ARB use No ACE-I/ARB use		se	
	(n=1336)	(n=4326)	SMD	(n=1179)	(n=4738)	SMD	
Age, median (IQR)	70 (60–79)	61 (55–73)	0.50	66 (58–77)	64 (56–78)	0.09	
Sex (male)	685 (51.3)	1830 (42.3)	0.18	525 (44.5)	2164 (45.7)	0.02	
Status as healthcare professional							
Physician	20 (1.5)	99 (2.3)	0.06	21 (1.8)	87 (1.8)	0.01	
Nurse	38 (2.8)	352 (8.1)	0.23	56 (4.7)	291 (6.1)	0.06	
Care assistent	39 (2.9)	223 (5.2)	0.11	39 (3.3)	193 (4.1)	0.04	
Use of antihypertensives (within prior 90 d	lays)						
ACE inhibitors	544 (40.7)	0	1.17	443 (37.6)	0	1.10	
Angiotensin receptor blockers	800 (59.9)	0	1.73	741 (62.9)	0	1.84	
Calcium channel blockers	376 (28.1)	213 (4.9)	0.66	157 (13.3)	771 (16.3)	0.08	
Thiazides	434 (32.5)	143 (3.3)	0.82	142 (12.0)	744 (15.7)	0.11	
Beta blockers	345 (25.8)	345 (8.0)	0.49	197 (16.7)	866 (18.3)	0.04	
Other antihypertensives	25 (1.9)	13 (0.3)	0.15	12 (1.1)	50 (1.1)	0.00	
Diagnoses (within prior 10 years)							
Hypertension	603 (45.1)	496 (11.5)	0.81	299 (25.4)	1277 (26.9)	0.04	
Atrial fibrillation	182 (13.6)	353 (8.2)	0.18	158 (13.4)	544 (11.5)	0.06	
Hospital-diagnosed obesity	134 (10.0)	256 (5.9)	0.15	91 (7.8)	436 (9.2)	0.05	
Angina	205 (15.3)	288 (6.7)	0.28	138 (11.7)	565 (11.9)	0.01	
Heart valve disease	90 (6.7)	113 (2.6)	0.20	51 (4.3)	176 (3.7)	0.03	
Alcohol abuse	47 (3.5)	125 (2.9)	0.04	47 (4.0)	211 (4.4)	0.02	
Diabetes	390 (29.2)	379 (8.8)	0.54	195 (16.6)	760 (16.0)	0.01	
Venous thromboembolism	57 (4.3)	171 (4.0)	0.02	56 (4.8)	192 (4.1)	0.04	
Dementia	89 (6.7)	263 (6.1)	0.02	91 (7.7)	285 (6.0)	0.07	
Myocardial infarction	73 (5.5)	99 (2.3)	0.17	45 (3.9)	147 (3.1)	0.04	
Liver disease	10 (0.7)	63 (1.5)	0.07	18 (1.6)	57 (1.2)	0.03	
Kidney disease	94 (7.0)	112 (2.6)	0.21	67 (5.7)	227 (4.8)	0.04	
Chronic pulmonary disease	158 (11.8)	391 (9.0)	0.09	145 (12.3)	488 (10.3)	0.06	
End-stage renal disease	(n<5)	11 (0.3)	0.02	13 (1.1)	11 (0.2)	0.10	
Heart failure	119 (8.9)	99 (2.3)	0.29	63 (5.4)	331 (7.0)	0.07	
Cancer	177 (13.2)	428 (9.9)	0.11	141 (12.0)	568 (12.0)	0.00	
Stroke	138 (10.3)	215 (5.0)	0.20	93 (7.9)	328 (6.9)	0.04	
Markers of smoking	430 (32.2)	1170 (27.0)	0.11	362 (30.7)	1321 (27.9)	0.06	
Medications (within prior 90 days)							
Statins	545 (40.8)	504 (11.7)	0.70	263 (22.3)	1145 (24.2)	0.04	
Antivirals	16 (1.2)	57 (1.3)	0.01	17 (1.4)	54 (1.1)	0.03	
Low-dose acetylsalicylic acid	208 (15.6)	199 (4.6)	0.37	112 (9.5)	395 (8.3)	0.04	
Vitamin K-antagonists	32 (2.4)	50 (1.2)	0.09	29 (2.4)	69 (1.4)	0.07	
Opioids	176 (13.2)	426 (9.8)	0.10	173 (14.7)	701 (14.8)	0.00	
Antidepressants	199 (14.9)	472 (10.9)	0.12	167 (14.1)	609 (12.9)	0.04	
Immunosuppresants	8 (0.6)	20 (0.5)	0.02	(n<5)	28 (0.6)	0.04	
Antipsychotics	52 (3.9)	138 (3.2)	0.04	40 (3.4)	169 (3.6)	0.01	
Glucocorticoids	55 (4.1)	150 (3.5)	0.03	61 (5.2)	168 (3.5)	0.08	
Hypnotics	89 (6.7)	218 (5.0)	0.07	95 (8.1)	359 (7.6)	0.02	
Loop diuretics	212 (15.9)	220 (5.1)	0.36	121 (10.3)	535 (11.3)	0.03	
Antibiotics	361 (27.0)	843 (19.5)	0.18	291 (24.7)	1106 (23.3)	0.03	

Continued

## Table 3 Continued

	Unweighted cohort			IPTW cohort		
	Current ACE-I/ARB use	No ACE-I/ARB use		Current ACE-I/ARB use	No ACE-I/ARB use	
	(n=1336)	(n=4326)	SMD	(n=1179)	(n=4738)	SMD
ocioeconomic factors						
Marital status						
Widowed	286 (21.4)	591 (13.7)	0.20	226 (19.1)	801 (16.9)	0.06
Divorced	195 (14.6)	745 (17.2)	0.07	194 (16.5)	734 (15.5)	0.03
Married	743 (55.6)	2519 (58.2)	0.05	641 (54.4)	2748 (58.0)	0.07
Unmarried	95 (7.1)	429 (9.9)	0.10	102 (8.7)	408 (8.6)	0.00
Unknown	17 (1.3)	42 (1.0)	0.03	15 (1.3)	46 (1.0)	0.03
Ethnicity						
Non-immigrant	1159 (86.8)	3665 (84.7)	0.06	1013 (86.0)	4008 (84.6)	0.04
First-generation or second-generation immigrant	177 (13.2)	661 (15.3)	0.06	166 (14.0)	730 (15.4)	0.04
Residence						
Non-urban	747 (55.9)	2397 (55.4)	0.01	644 (54.6)	2685 (56.7)	0.04
Urban	589 (44.1)	1929 (44.6)	0.01	535 (45.4)	2052 (43.3)	0.04

ACE-I, ACE inhibitors; ARB, angiotensin receptor blocker; SMD, standardised mean difference.

	Exposed		Unexposed		Comparison		
Outcome	Events	Risk (%)	Events	Risk (%)	Risk difference (%)	Relative risk	
Current ACE-I/ARB use	ers compared with	current CCB users					
Unweighted cohort							
Death	126/1065	11.8 (9.9–13.8)	43/289	14.9 (10.8–19.0)	-3.0 (-7.6-1.5)	0.80 (0.58–1.10)	
ICU admission	66/1059	6.2 (4.8–7.7)	16/288	5.6 (2.9–8.2)	0.7 (-2.3-3.7)	1.12 (0.66–1.91)	
Mechanical ventilation	51/1065	4.8 (3.5–6.1)	16/289	5.5 (2.9–8.2)	-0.7 (-3.7-2.2)	0.86 (0.50–1.49)	
Hospital admission	361/1001	36.1 (33.1–39.0)	125/273	45.8 (39.9–51.7)	-9.7 (-16.33.1)	0.79 (0.68–0.92)	
Acute dialysis	NA	NA	NA	NA	0.2 (-1.4-1.8)	1.13 (0.38–3.35)	
IPTW cohort							
Death	131/1065	12.3 (10.3–14.4)	40/285	13.9 (9.2–18.6)	-1.5 (-6.7-3.6)	0.89 (0.61–1.30)	
ICU admission	66/1058	6.3 (4.8–7.8)	15/283	5.4 (2.3–8.4)	0.9 (-2.4-4.3)	1.17 (0.64–2.16)	
Mechanical ventilation	52/1065	4.8 (3.5–6.2)	17/285	5.9 (2.6–9.2)	-1.1 (-4.6-2.5)	0.82 (0.44–1.53)	
Hospital admission	366/1001	36.5 (33.5–39.6)	118/274	43.3 (36.0–50.6)	-6.7 (-14.7-1.2)	0.84 (0.70–1.02)	
Acute dialysis	NA	NA	NA	NA	0.7 (-0.7-2.0)	1.62 (0.53–4.92)	
Current ace-i/ARB use	rs compared with a	ace-i/ARB non-users aged	50 years or older				
Unweighted cohort							
Death	166/1336	12.4 (10.7–14.2)	367/4326	8.5 (7.7–9.3)	3.9 (2.0–5.9)	1.46 (1.23–1.74)	
ICU admission	104/1329	7.8 (6.4–9.3)	166/4321	3.8 (3.3–4.4)	4.0 (2.4–5.5)	2.04 (1.61–2.58)	
Mechanical ventilation	81/1336	6.1 (4.8–7.3)	121/4326	2.8 (2.3–3.3)	3.3 (1.9–4.6)	2.17 (1.65–2.85)	
Hospital admission	507/1258	40.3 (37.6–43.0)	1097/4192	26.2 (24.8–27.5)	14.1 (11.1–17.2)	1.54 (1.42–1.68)	
Acute dialysis	27/1325	2.0 (1.3–2.8)	37/4309	0.9 (0.6–1.1)	1.2 (0.4–2.0)	2.37 (1.45–3.88)	
IPTW cohort							
Death	112/1179	9.5 (7.5–11.5)	512/4738	10.8 (9.1–12.5)	-1.3 (-3.9-1.3)	0.88 (0.68–1.14)	
ICU admission	75/1173	6.4 (4.7–8.1)	243/4733	5.1 (2.2–8.1)	1.3 (-2.1-4.7)	1.25 (0.66–2.36)	
Mechanical ventilation	62/1179	5.3 (3.7–6.9)	201/4738	4.2 (1.3–7.2)	1.0 (-2.3-4.4)	1.24 (0.58–2.66)	
Hospital admission	362/1118	32.3 (28.8–35.9)	1505/4546	33.1 (29.5–36.7)	-0.8 (-5.8-4.3)	0.98 (0.84–1.14)	
Acute dialysis	27/1172	2.3 (1.1–3.5)	36/4710	0.8 (0.5–1.0)	1.6 (0.3–2.8)	3.04 (1.64–5.65)	

ACE-I, ACE inhibitors; ARB, angiotensin receptor blocker; CCB, calcium channel blocker; ICU, intensive care unit; IPTW, inverse probability of treatment weighting; NA, not available.

# **Respiratory infection**

departments or during hospitalisation at New York University Langone Health.<sup>6</sup> They examined use of ACE-Is/ARBs and other antihypertensive medications and a positive SARS-CoV-2 test result, as well as severity of COVID-19, and found no substantial association. The study was limited by its inclusion of only 18 months of chronic disease history, as well as by defining ACE-I/ ARB exposure as any treatment yes/no within 18 months before the SARS-CoV-2 tests without evidence of discontinuation in the previous 30 days. In addition, severity of illness (defined by ICU admission, MV or death) was assessed only at one time point at the end of the study period, that is, length of follow-up differed by date of inclusion.<sup>6</sup> An Italian case-control study included 6272 patients diagnosed with COVID-19 matched to 30 759 controls. The authors found that neither ACE-I nor ARB use was associated with risk of and severity of COVID-19.7 Exposure was defined as redemption of any prescription in 2019. A Danish study of 4480 patients with hospital-diagnosed COVID-19, found no clear association between ACE-I/ARB use and 30-day mortality (HR 0.83, 95% CI 0.67 to 1.03).<sup>13</sup>

#### **Potential mechanisms**

We have no data to examine whether conflicting mechanisms produced the neutral association with SARS-CoV-2 infection and outcomes that we observed. Cell-surface ACE2 expression, which is likely increased in patients treated with ACE-Is/ARBs, may facilitate SARS-CoV-2 entry into cells.<sup>5</sup> However, ACE2 expression also has been shown to protect against development of severe acute lung injury in other infectious diseases.<sup>1 4 35</sup> The null findings of our case-control risk analyses argue against any clinically relevant effect of ACE-I/ARB-mediated ACE2 upregulation in increasing SARS-CoV-2 risk in general. These issues will be addressed in several planned trials on the impact of losartan on organ dysfunction and mortality in patients hospitalised with SARS-CoV-2 infection (ClinicalTrials.gov Identifier: NCT04312009 and NCT04328012).

#### CONCLUSION

Our study showed that ACE-I/ARB use was not associated with a consistently higher rate of acquiring a positive SARS-CoV-2 test among tested individuals. Further, ACE-I/ARB use was not associated with increased mortality and other adverse outcomes among patients with microbiologically confirmed SARS-CoV-2 infection. These results support the recommendations of continuing treatment with ACE-Is/ARBs during the COVID-19 pandemic.

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### REFERENCES

- Sanders JM, Monogue ML, Jodlowski TZ, et al. Pharmacologic treatments for coronavirus disease 2019 (COVID-19): a review. JAMA 2020;323:1824–36.
- 2 Fang L, Karakiulakis G, Roth M. Are patients with hypertension and diabetes mellitus at increased risk for COVID-19 infection? *Lancet Respir Med* 2020;8:e21.
- 3 Sommerstein R, Gräni C. Rapid response: re: preventing a covid-19 pandemic: ACE inhibitors as a potential risk factor for fatal Covid-19. *BMJ* 2020;368:m810 https:// www.bmj.com/content/368/bmj.m810/rr-2
- 4 Patel AB, Verma A. COVID-19 and angiotensin-converting enzyme inhibitors and angiotensin receptor blockers: what is the evidence? JAMA 2020;323:1769–70.
- 5 Hoffmann M, Kleine-Weber H, Schroeder S, *et al.* SARS-CoV-2 cell entry depends on ACE2 and TMPRSS2 and is blocked by a clinically proven protease inhibitor. *Cell* 2020;181:271–80.
- 6 Reynolds HR, Adhikari S, Pulgarin C, et al. Renin-Angiotensin-Aldosterone system inhibitors and risk of Covid-19. N Engl J Med 2020;382:2441–8.
- 7 Mancia G, Rea F, Ludergnani M, *et al.* Renin-Angiotensin-Aldosterone system blockers and the risk of Covid-19. *N Engl J Med* 2020;382:2431–40.
- 8 Mehta N, Kalra A, Nowacki AS, et al. Association of use of angiotensin-converting enzyme inhibitors and angiotensin II receptor blockers with testing positive for coronavirus disease 2019 (COVID-19). JAMA Cardiol 2020;5:1020–6.
- 9 Holt A, Mizrak I, Lamberts M, et al. Influence of inhibitors of the renin-angiotensin system on risk of acute respiratory distress syndrome in Danish hospitalized COVID-19 patients. J Hypertens 2020;38:1612–3.

- 10 Zhang P, Zhu L, Cai J, et al. Association of inpatient use of angiotensin-converting enzyme inhibitors and angiotensin II receptor blockers with mortality among patients with hypertension hospitalized with COVID-19. Circ Res 2020;126:1671–81.
- 11 Li J, Wang X, Chen J, et al. Association of renin-angiotensin system inhibitors with severity or risk of death in patients with hypertension hospitalized for coronavirus disease 2019 (COVID-19) infection in Wuhan, China. JAMA Cardiol 2020;5:825–30.
- 12 de Abajo FJ, Rodríguez-Martín S, Lerma V, et al. Use of renin-angiotensin-aldosterone system inhibitors and risk of COVID-19 requiring admission to hospital: a casepopulation study. *The Lancet* 2020;395:1705–14.
- 13 Fosbøl EL, Butt JH, Østergaard L, et al. Association of angiotensin-converting enzyme inhibitor or angiotensin receptor blocker use with COVID-19 diagnosis and mortality. JAMA 2020;324:168–77.
- 14 Felice C, Nardin C, Di Tanna GL, et al. Use of RAAS inhibitors and risk of clinical deterioration in COVID-19: results from an Italian cohort of 133 hypertensives. Am J Hypertens 2020;33:944–8.
- 15 Gao C, Cai Y, Zhang K, et al. Association of hypertension and antihypertensive treatment with COVID-19 mortality: a retrospective observational study. Eur Heart J 2020;41:2058–66.
- 16 Gnavi R, Demaria M, Picariello R, *et al.* Therapy with agents acting on the renin-angiotensin system and risk of SARS-CoV-2 infection. *Clin Infect Dis* 2020. doi:10.1093/cid/ciaa634. [Epub ahead of print: 22 May 2020].
- 17 European Medicines Agency. EMA advises continued use of medicines for hypertension, heart or kidney disease during COVID-19 pandemic. Available: https://www.ema.europa.eu/en/news/ema-advises-continued-use-medicineshypertension-heart-kidney-disease-during-covid-19-pandemic [Accessed 13 Oct 2020].
- 18 Kuster GM, Pfister O, Burkard T, et al. SARS-CoV2: should inhibitors of the renin–angiotensin system be withdrawn in patients with COVID-19? Eur Heart J 2020;41:1801–3.
- 19 Schmidt M, Schmidt SAJ, Adelborg K, et al. The Danish health care system and epidemiological research: from health care contacts to database records. Clin Epidemiol 2019;11:563–91.
- 20 Reilev M, Kristensen KB, Pottegård A, et al. Characteristics and predictors of hospitalization and death in the first 11 122 cases with a positive RT-PCR test for SARS-CoV-2 in Denmark: a nationwide cohort. Int J Epidemiol 2020. doi:10.1093/ije/ dyaa140. [Epub ahead of print: 05 Sep 2020].

- 21 Voldstedlund M, Haarh M, Mølbak K, *et al*. The Danish microbiology database (MiBa) 2010 to 2013. *Euro Surveill* 2014;19:20667.
- 22 Schmidt M, Schmidt SAJ, Sandegaard JL, et al. The Danish national patient registry: a review of content, data quality, and research potential. *Clin Epidemiol* 2015;7:449–90.
- 23 Schmidt M, Pedersen L, Sørensen HT. The Danish civil registration system as a tool in epidemiology. *Eur J Epidemiol* 2014;29:541–9.
- 24 Pottegård A, Schmidt SAJ, Wallach-Kildemoes H, *et al*. Data resource profile: the Danish national prescription registry. *Int J Epidemiol* 2017;46:798–798f.
- 25 Helweg-Larsen K. The Danish register of causes of death. Scand J Public Health 2011;39:26–9.
- 26 Blichert-Hansen L, Nielsson MS, Nielsen RB, et al. Validity of the coding for intensive care admission, mechanical ventilation, and acute dialysis in the Danish national patient registry: a short report. Clin Epidemiol 2013;5:9–12.
- 27 Christiansen CF, Møller MH, Nielsen H, *et al.* The Danish intensive care database. *Clin Epidemiol* 2016;8:525–30.
- 28 Brookhart MA, Wyss R, Layton JB, et al. Propensity score methods for confounding control in nonexperimental research. *Circ Cardiovasc Qual Outcomes* 2013;6:604–11.
- 29 van Kasteren PB, van der Veer B, van den Brink S, *et al*. Comparison of seven commercial RT-PCR diagnostic kits for COVID-19. *J Clin Virol* 2020;128:104412.
- 30 Richardson S, Hirsch JS, Narasimhan M, *et al*. Presenting characteristics, comorbidities, and outcomes among 5700 patients hospitalized with COVID-19 in the new York City area. *JAMA* 2020;323:2052–9.
- 31 Pottegård A, Kurz X, Moore N, *et al*. Considerations for pharmacoepidemiological analyses in the SARS-CoV-2 pandemic. *Pharmacoepidemiol Drug Saf* 2020;29:825–31.
- 32 Phua J, Weng L, Ling L, *et al*. Intensive care management of coronavirus disease 2019 (COVID-19): challenges and recommendations. *Lancet Respir Med* 2020;8:506–17.
- 33 Sundbøll J, Adelborg K, Munch T, et al. Positive predictive value of cardiovascular diagnoses in the Danish national patient registry: a validation study. BMJ Open 2016;6:e012832.
- 34 Mackey K, King VJ, Gurley S, et al. Risks and impact of angiotensin-converting enzyme inhibitors or angiotensin-receptor blockers on SARS-CoV-2 infection in adults: a living systematic review. Ann Intern Med 2020;173:195–203.
- 35 Vaduganathan M, Vardeny O, Michel T, *et al*. Renin-Angiotensin-Aldosterone system inhibitors in patients with Covid-19. *N Engl J Med* 2020;382:1653–9.